

Figure 7.2.1. Transformation method for generating a random deviate y from a known probability distribution p(y). The indefinite integral of p(y) must be known and invertible. A uniform deviate x is chosen between 0 and 1. Its corresponding y on the definite-integral curve is the desired deviate.

Let's see what is involved in using the above *transformation method* to generate some arbitrary desired distribution of y's, say one with p(y) = f(y) for some positive function f whose integral is 1. (See Figure 7.2.1.) According to (7.2.4), we need to solve the differential equation

$$\frac{dx}{dy} = f(y) \tag{7.2.6}$$

But the solution of this is just x = F(y), where F(y) is the indefinite integral of f(y). The desired transformation which takes a uniform deviate into one distributed as f(y) is therefore

$$y(x) = F^{-1}(x) \tag{7.2.7}$$

where F^{-1} is the inverse function to *F*. Whether (7.2.7) is feasible to implement depends on whether the *inverse function of the integral of f*(*y*) is itself feasible to compute, either analytically or numerically. Sometimes it is, and sometimes it isn't.

Incidentally, (7.2.7) has an immediate geometric interpretation: Since F(y) is the area under the probability curve to the left of y, (7.2.7) is just the prescription: choose a uniform random x, then find the value y that has that fraction x of probability area to its left, and return the value y.

Normal (Gaussian) Deviates

Transformation methods generalize to more than one dimension. If x_1, x_2 , ... are random deviates with a *joint* probability distribution $p(x_1, x_2, ...)$ $dx_1 dx_2 ...,$ and if $y_1, y_2, ...$ are each functions of all the x's (same number of y's as x's), then the joint probability distribution of the y's is

$$p(y_1, y_2, \ldots) dy_1 dy_2 \ldots = p(x_1, x_2, \ldots) \left| \frac{\partial(x_1, x_2, \ldots)}{\partial(y_1, y_2, \ldots)} \right| dy_1 dy_2 \ldots$$
(7.2.8)

where $|\partial(-)/\partial(-)|$ is the Jacobian determinant of the x's with respect to the y's (or reciprocal of the Jacobian determinant of the y's with respect to the x's).

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An important example of the use of (7.2.8) is the *Box-Muller* method for generating random deviates with a normal (Gaussian) distribution,

$$p(y)dy = \frac{1}{\sqrt{2\pi}}e^{-y^2/2}dy$$
(7.2.9)

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Consider the transformation between two uniform deviates on (0,1), x_1, x_2 , and two quantities y_1, y_2 ,

$$y_1 = \sqrt{-2\ln x_1} \cos 2\pi x_2$$

$$y_2 = \sqrt{-2\ln x_1} \sin 2\pi x_2$$
(7.2.10)

Equivalently we can write

$$x_{1} = \exp\left[-\frac{1}{2}(y_{1}^{2} + y_{2}^{2})\right]$$

$$x_{2} = \frac{1}{2\pi}\arctan\frac{y_{2}}{y_{1}}$$
(7.2.11)

Now the Jacobian determinant can readily be calculated (try it!):

$$\frac{\partial(x_1, x_2)}{\partial(y_1, y_2)} = \begin{vmatrix} \frac{\partial x_1}{\partial y_1} & \frac{\partial x_1}{\partial y_2} \\ \frac{\partial x_2}{\partial y_1} & \frac{\partial x_2}{\partial y_2} \end{vmatrix} = -\left[\frac{1}{\sqrt{2\pi}}e^{-y_1^2/2}\right] \left[\frac{1}{\sqrt{2\pi}}e^{-y_2^2/2}\right]$$
(7.2.12)

Since this is the product of a function of y_2 alone and a function of y_1 alone, we see that each y is independently distributed according to the normal distribution (7.2.9).

One further trick is useful in applying (7.2.10). Suppose that, instead of picking uniform deviates x_1 and x_2 in the unit square, we instead pick v_1 and v_2 as the ordinate and abscissa of a random point inside the unit circle around the origin. Then the sum of their squares, $R^2 \equiv v_1^2 + v_2^2$ is a uniform deviate, which can be used for x_1 , while the angle that (v_1, v_2) defines with respect to the v_1 axis can serve as the random angle $2\pi x_2$. What's the advantage? It's that the cosine and sine in (7.2.10) can now be written as $v_1/\sqrt{R^2}$ and $v_2/\sqrt{R^2}$, obviating the trigonometric function calls!

We thus have

#include <math.h>

float gasdev(long *idum)
Returns a normally distributed deviate with zero mean and unit variance, using ran1(idum)
as the source of uniform deviates.
f

```
float ran1(long *idum);
static int iset=0;
static float gset;
float fac,rsq,v1,v2;
if (*idum < 0) iset=0;
if (iset == 0) {
    do {
        v1=2.0*ran1(idum)-1.0;
        v2=2.0*ran1(idum)-1.0;
        rsq=v1*v1+v2*v2;
    }
}</pre>
```

Reinitialize. We don't have an extra deviate handy, so

pick two uniform numbers in the square extending from -1 to +1 in each direction, see if they are in the unit circle,

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```
} while (rsq >= 1.0 || rsq == 0.0);
                                                   and if they are not, try again.
    fac=sqrt(-2.0*log(rsq)/rsq);
    Now make the Box-Muller transformation to get two normal deviates. Return one and
    save the other for next time.
    gset=v1*fac;
    iset=1;
                                           Set flag.
    return v2*fac;
} else {
                                           We have an extra deviate handy,
    iset=0;
                                           so unset the flag,
    return gset;
                                           and return it
}
```

See Devroye [1] and Bratley [2] for many additional algorithms.

CITED REFERENCES AND FURTHER READING:

Devroye, L. 1986, Non-Uniform Random Variate Generation (New York: Springer-Verlag), §9.1. [1]

- Bratley, P., Fox, B.L., and Schrage, E.L. 1983, *A Guide to Simulation* (New York: Springer-Verlag). [2]
- Knuth, D.E. 1981, Seminumerical Algorithms, 2nd ed., vol. 2 of The Art of Computer Programming (Reading, MA: Addison-Wesley), pp. 116ff.

7.3 Rejection Method: Gamma, Poisson, Binomial Deviates

The rejection method is a powerful, general technique for generating random deviates whose distribution function p(x)dx (probability of a value occurring between x and x + dx) is known and computable. The rejection method does not require that the cumulative distribution function [indefinite integral of p(x)] be readily computable, much less the inverse of that function — which was required for the transformation method in the previous section.

The rejection method is based on a simple geometrical argument:

Draw a graph of the probability distribution p(x) that you wish to generate, so that the area under the curve in any range of x corresponds to the desired probability of generating an x in that range. If we had some way of choosing a random point *in two dimensions*, with uniform probability in the *area* under your curve, then the x value of that random point would have the desired distribution.

Now, on the same graph, draw any other curve f(x) which has finite (not infinite) area and lies everywhere *above* your original probability distribution. (This is always possible, because your original curve encloses only unit area, by definition of probability.) We will call this f(x) the *comparison function*. Imagine now that you have some way of choosing a random point in two dimensions that is uniform in the area under the comparison function. Whenever that point lies outside the area under the original probability distribution, we will *reject* it and choose another random point. Whenever it lies inside the area under the original probability distribution, we will *accept* it. It should be obvious that the accepted points are uniform in the accepted area, so that their x values have the desired distribution. It

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